

Market Externalities of Large Unemployment Insurance Extensions

Rafael Lalive, Camille Landais & Josef Zweimuller

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Motivation:

What is the effect of increasing generosity of UI on labor market outcomes?

- We \approx know what micro effect is
 - ▶ In *theory*, increase in UI unambiguously increase U duration
 - ▶ *Empirically*, large number of well-identified micro estimates
- What about macro effect?
 - ▶ In *theory*, large literature on equilibrium search & matching, but anything goes regarding externalities
 - ▶ *Empirically*, difficulty of estimating G-E effects of UI and to analyze how micro and macro estimates differ

UI and labor market externalities:

- Market externality:
Whenever (UI induced) variations in the search effort of some unemployed affect job finding probability of other unemployed in the same labor market
- Market externality \neq incidence:
In market with frictions, efficiency is usually not achieved, so that (UI induced) variations in behaviors have first order welfare effects

This paper:

- Regional Extended Benefit Program (REBP): Large extensions of UI in Austria
 - ▶ Unique quasi-experimental setting to identify market externalities
 - ▶ Strong evidence of positive effects of REBP on untreated workers in treated labor markets
- Discuss how evidence relates to different search & matching models:
 - ▶ Evidence refutes predictions of Nash bargaining / flexible wage models
 - ▶ Evidence in line with job-rationing models

Related literature:

- Theoretical literature on pecuniary externalities:
 - ▶ Geanakoplos & Polemarchakis (1986), etc.
- Literature on optimal UI:
 - ▶ Direct continuity of LMS (2012)
- Empirical literature on identification of spillovers of policy interventions
 - ▶ General literature on spillovers: Duflo & Saez (2003)
 - ▶ Spillovers of active labor market policies: Crepon & al. (2012), Ferracci & al. (2010), Blundell, & al. (2004).
 - ▶ Spillovers of UI: Levine (1993)

- ① Introduction
- ② Conceptual framework
- ③ Institutional background
- ④ Empirical strategy
- ⑤ Results
- ⑥ Calibrations

Labor Market with Matching Frictions

- u unemployed workers:
 - ▶ Exert search effort e
 - ▶ e function of wedge in consumption $\Delta c = c^e - c^u$
- v vacancies.
- Number of matches: $m(e \cdot u, v) = \omega_m \cdot (e \cdot u)^\eta \cdot v^{1-\eta}$
- Labor market tightness: $\theta \equiv v / (e \cdot u)$
- Job-finding proba: $e \cdot f(\theta) = e \cdot m(1, \theta)$.
- Vacancy-filling proba: $q(\theta) = m(1/\theta, 1)$.

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Labor market equilibrium

- Aggregate labor supply (from equality of in- and outflows into employment):

$$n^s(e(\theta, \Delta c), \theta)$$

- Aggregate labor demand (from firm's maximisation program):

$$n^d(\theta)$$

- Labor market equilibrium:

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Figure 1 : Externalities in a model with Nash bargaining

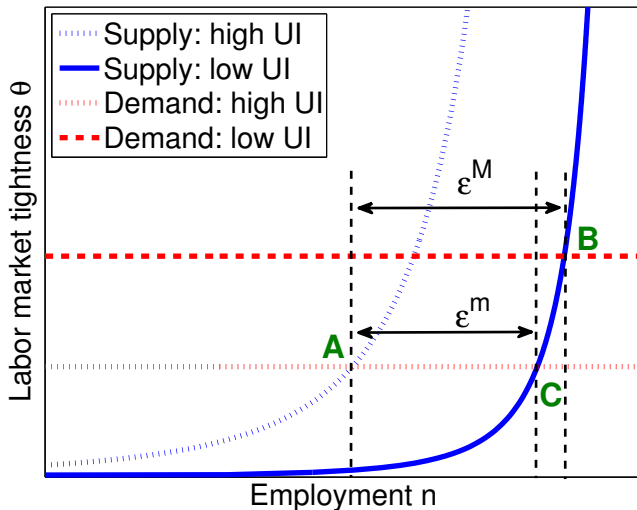
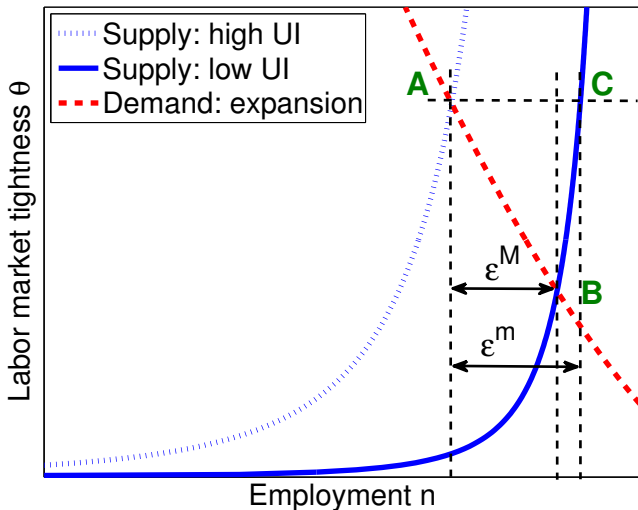


Figure 2 : Labor market equilibrium in a Michailat model



Externalities in different matching models

- In models with flexible wages:
 - ▶ $\downarrow \Delta c \Rightarrow \uparrow w \Rightarrow \downarrow n^d$
 - ▶ Macro effect larger than micro effect
- In models with rigid wages & diminishing returns:
 - ▶ $\downarrow \Delta c \Rightarrow \uparrow (f' - w) \Rightarrow \uparrow n^d$
 - ▶ Macro effect smaller than micro effect

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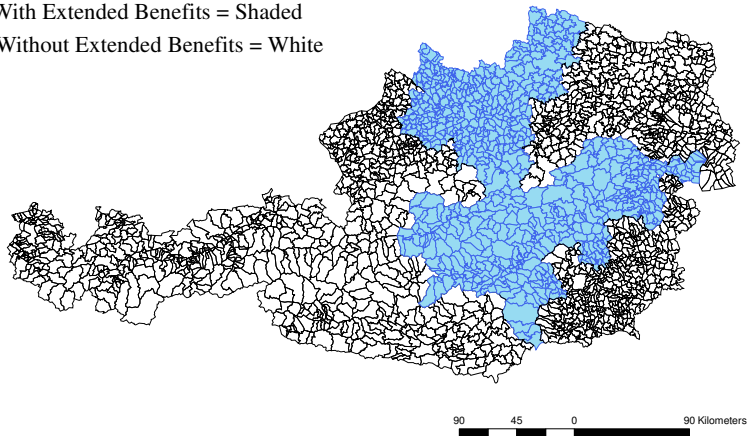
REBP reform in Austria

- Large UI benefit extension program enacted in Austria
 - ▶ 209 weeks instead of 52 weeks
- Eligibility requirements:
 - ▶ Age: more than 50
 - ▶ Reside in selected regions at least 6 months before becoming unemployed
 - ▶ At least 15 years of continuous work history in the past 25 years
 - ▶ Spell beginning between June 1988 and Dec 1993

Figure 3 : Austrian regions by REBP treatment status

With Extended Benefits = Shaded

Without Extended Benefits = White



Data:

- Universe of UI spells in Austria from 1980 to 2010:
 - ▶ Info on age, residence, education, marital status, etc...
- Universe of social security data in Austria from 1949 to 2010:
 - ▶ Info on each employment spell
 - ▶ Compute experience in past 25 years
 - ▶ Merge with UI data to determine REBP eligibility
 - ▶ Info on wages, industry, tenure,

Sample selection:

- Endogeneity of choice of REBP regions:
 - ▶ Regions are not selected at random: restructuring of steel sector
 - ▶ Remove all steel sector workers (at most 15% of unemployed in treated regions), and all workers in related industries
- Early retirement programs:
 - ▶ Women can go directly from REBP to early retirement programs
 - ▶ We focus only on men 50 to 54 bc they cannot go directly from REBP to early retirement

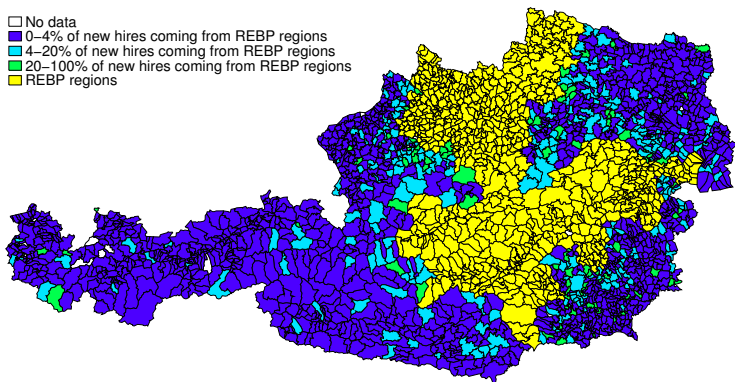
Empirical strategy:

- **First stage:** Compare treated workers in treated regions and untreated regions before/during/after
- **Second stage:** Compare untreated workers in treated and untreated regions before/during/after
- Identification assumptions:
 - ▶ Treated and untreated regions are somehow isolated
 - ▶ Unobserved differences between treated and untreated workers fixed over time
 - ▶ Unobserved differences between labor markets are fixed over time

Table 1 : SUMMARY STATISTICS:

| | (1) | (2) | (3) | (4) |
|--|----------|----------|------------|---------|
| A. All workers | | | | |
| treated vs untreated counties before 1988 | | | | |
| | M=0 | M=1 | Difference | p-value |
| Age | 51.9 | 51.9 | 0 | .366 |
| U duration | 18.7 | 19.4 | -.7 | .12 |
| Non employment duration | 31.7 | 29.9 | 1.8 | .018 |
| Fraction spells > 100 wks | .033 | .039 | -.006 | .023 |
| Fraction spells > 26 wks | .135 | .122 | .013 | .016 |
| Real wage before spell | 52.1 | 50.5 | 1.6 | 0 |
| Real wage after spell | 51.8 | 50.8 | 1.1 | 0 |
| White Collar | .063 | .035 | .028 | 0 |
| Fraction not in construction | .38 | .369 | .011 | .148 |
| B. Treated workers vs untreated workers | | | | |
| in treated counties before 1988 | | | | |
| | T=0 | T=1 | Difference | p-value |
| Age | 51.8 | 51.9 | -.1 | .181 |
| Experience | 4089.365 | 8292.634 | -4203.269 | 0 |
| U duration | 16.3 | 19.6 | -3.3 | .025 |
| Non employment duration | 52.5 | 28 | 24.5 | 0 |
| Fraction spells > 100 wks | .018 | .041 | -.023 | .022 |
| Fraction spells > 26 wks | .091 | .124 | -.033 | .056 |
| Real wage before spell | 47.3 | 50.8 | -3.6 | 0 |
| Real wag after spell | 47.4 | 51 | -3.6 | 0 |
| White Collar | .01 | .037 | -.027 | .006 |
| Fraction not in construction | .345 | .371 | -.026 | .307 |

Figure 4 : Local labor markets integration: Fraction of new hires from REBP regions in total number of new hires by county



Sample: male age 50 to 54 in non steel-related industries, 1980-1987.

Figure 5 : Difference in U duration between REBP and non REBP regions: male 50-54 with more than 15 years of experience

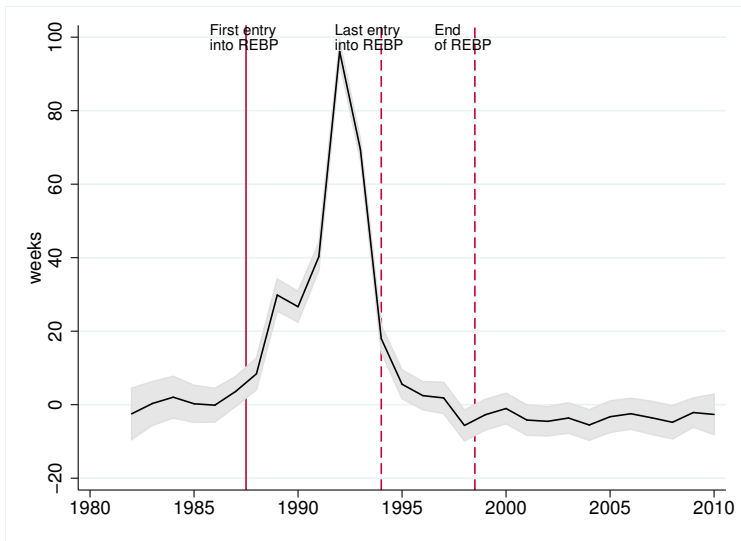


Figure 6 : Difference in U duration between REBP and non REBP regions: male 50-54 with less than 15 years of experience



Baseline specifications:

$$Y_{irt} = \alpha + \overbrace{\beta_0 \cdot Z_{irt} \cdot R_r \cdot T_t}^{\text{Effect of REBP on treated}} + \overbrace{\gamma_0 \cdot (1 - Z_{irt}) \cdot R_r \cdot T_t}^{\text{Effect of REBP on non-treated}} \\ + \eta_0 R_r + \eta_1 B_{irt} + \eta_2 B_{irt} \cdot R_r \\ + \sum \nu_t + \sum \eta_3 B_{irt} \cdot \iota_t + X'_{it} \rho + \varepsilon_{irt}$$

- R_r : indicator for residing in REBP region
- T_t : indicator for spell starting btw June 1988 and Dec 1997
- $B_{irt} = \mathbb{1}[exp > 15]$: indicator for more than 15 yrs of exp
- $Z_{irt} = B_{irt} \cdot \tilde{T}_t$: indicator for being eligible to REBP extensions

Table 2 : Baseline estimates of the treatment effect of REBP on treated unemployed and untreated unemployed

| | (1) | (2) Unemployment duration | (3) | (4) | (5) Non-empl. duration | (6) Spell >100 wks | (7) Spell >26 wks |
|---|----------------------|------------------------------|----------------------|----------------------|------------------------------|--------------------------|-------------------------|
| β_0 | 62.41*** (9.565) | 54.57*** (8.345) | 55.48*** (9.051) | 58.14*** (9.159) | 26.03*** (5.797) | 0.233*** (0.0312) | 0.236*** (0.0290) |
| γ_0 | -6.941*** (1.690) | -7.165*** (2.017) | -11.86*** (1.640) | -8.979*** (1.433) | -9.725*** (1.487) | -0.0186*** (0.00509) | -0.0297** (0.0116) |
| Educ., marital status, industry, citizenship | | × | × | × | × | × | × |
| Preexisting trends | | | | | | | |
| by region | | | × | | | | |
| by region \times exp | | | | × | × | × | × |
| N | 127802 | 126091 | 126091 | 126091 | 106164 | 126091 | 126091 |

S.e. clustered at the year \times region level in parentheses. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.010$.

Potential confounders:

■ Confounder 1: **selection**

- ▶ Self-selection into unemployment affected by the reform for non-treated group in treated counties
- ▶ If anything, bias likely to attenuate estimate of spillover effect on non-treated

■ Confounder 2: **differential region-specific shocks**

- ▶ REBP regions experience positive shock on labor market conditions at the time REBP was implemented
- ▶ If anything, we expect negative shock if REBP regions endogenously selected

Table 3 : Testing for selection: inflow rate into unemployment and log real wage in previous job

| | (1) log separation rate | (2) log real wage in previous job | (3) log real wage in previous job |
|--------------|-------------------------------|---|---|
| eligible | 0.287*** (0.0355) | | |
| non-eligible | -0.0346 (0.0306) | | |
| β_0 | | 0.144** (0.0691) | 0.132** (0.0614) |
| γ_0 | | -0.0638 (0.0629) | -0.0479 (0.0608) |
| N | 1733 | 114770 | 112242 |

Standard errors in parentheses, * $p < 0.10$, ** $p < 0.05$, *** $p < 0.010$

Table 4 : Using regions close to REBP border with high labor market integration as spillover group

| | (1) Unemployment duration | (2) | (3) | (4) Non-empl. duration | (5) Spell >100 wks | (6) Spell >26 wks |
|---|------------------------------|---------------------|---------------------|------------------------------|--------------------------|-------------------------|
| β_0 | 66.20*** (10.13) | 58.24*** (8.865) | 65.09*** (9.869) | 27.68*** (6.298) | 0.254*** (0.0339) | 0.251*** (0.0316) |
| γ_0 | -1.813 (3.323) | -1.588 (2.954) | -3.110 (3.261) | -3.446 (2.563) | -0.0117 (0.0118) | -0.0602** (0.0257) |
| Educ., marital status, industry, citizenship | | × | × | × | × | × |
| Preexisting trends by region | | | × | × | × | × |
| <i>N</i> | 160714 | 157578 | 159104 | 135702 | 159104 | 159104 |

S.e. clustered at the year×region level in parentheses. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.010$

Table 5 : Effects of REBP on subsequent wages and match quality

| | (1) | (2) | (3) | (4) | (5) | (6) |
|---|------------------------------|-----------------------|---|--------------------|----------------------------------|-------------------|
| | log real wage in next job | | wage drop from next to previous job | | distance to next job (min) | |
| β_0 | -0.0236 (0.0154) | -0.0381** (0.0152) | -0.157 (0.214) | -0.0904 (0.208) | -0.456 (0.554) | 0.223 (0.549) |
| γ_0 | 0.00515 (0.0448) | -0.0477 (0.0441) | 0.269 (0.591) | 0.462 (0.562) | -0.233 (1.138) | 2.476* (1.240) |
| Educ., marital status, industry, citizenship | | × | | × | | × |
| N | 90345 | 88634 | 94503 | 92719 | 103678 | 101715 |

Standard errors in parentheses. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.010$

Conclusion:

- Identification of positive effects of increasing UI on untreated workers in the same labor market
- Externalities matter in the labor market and must be taken into account for optimal UI
- Next steps: heterogeneity analysis

Figure 7 : Local labor markets integration: Fraction of new hires from non-REBP regions in total number of new hires by county

